

Boolean formula f on n variables x_1, \dots, x_n

CNF = conjunctive normal form

conjunction (AND) of clauses \wedge

each clause is the disjunction of literals
(OR)

DNF = disjunctive normal form

OR of clauses \wedge each clause is the AND of literals

(So need to satisfy all literals in ≥ 1 clause)

Easy to find a satisfying assignment for a formula f in DNF.

But how many satisfying assignments?

#DNF:

input: formula f in DNF

with n variables x_1, \dots, x_n & m clauses C_1, \dots, C_m .

output: the # of satisfying assignments.

#P-complete to compute exactly for all f in time $\text{Poly}(n, m)$.

Can we approximate it?

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Strongest form of approximation:

For input f , let $N(f) = \#$ of satisfying assignments.

FPRAS = fully polynomial randomized approximation scheme.

Input: formula f , accuracy $\epsilon > 0$, error $\delta > 0$.

Algorithm is an FPRAS if it outputs OUT s.t.

$$\Pr(\text{OUT}(1-\epsilon) \leq N(f) \leq \text{OUT}(1+\epsilon)) \geq 1 - \delta$$

in time $\text{Poly}(n, \frac{1}{\epsilon}, \log(\frac{1}{\delta}))$.

Suffices to do so with:

$$\Pr(\text{OUT}(1-\epsilon) \leq N(f) \leq \text{OUT}(1+\epsilon)) \geq 3/4$$

then can run $O(\log(\frac{1}{\delta}))$ trials,

& take the median of their outputs.

By Chernoff bounds, this is an FPRAS.

(3)

Monte Carlo approach:

Set S that we want to estimate $|S|$.

Find trivial set Σ where $\Sigma \supset S$ & $|\Sigma|$ is known.

Sample from Σ & look at prob. that it lies in S .

Let X_1, \dots, X_t be uniform, random samples from Σ ,

& let $y_i = \begin{cases} 1 & \text{if } X_i \in S \\ 0 & \text{if not} \end{cases}$

$$\text{Note, } E[Y_i] = \Pr(X_i \in S) = \frac{|S|}{|\Sigma|} := \mu$$

Look at $Y = \frac{1}{t} \sum_i Y_i$

$$\text{Note, } |S| = \mu |\Sigma|$$

$$\text{Let } \hat{Y} = |\Sigma| Y = \frac{|\Sigma|}{t} \sum_i Y_i$$

How large of t to get a good estimate of $|S|$?

(4)

Apply Chernoff bounds:

$$\begin{aligned} & \Pr\left(\left|\frac{1}{|S|} \sum_i Y_i - \frac{1}{|S|}\right| \geq \epsilon \frac{|S|}{|S|}\right) \\ &= \Pr\left(\left|\frac{1}{|S|} \sum_i Y_i - \frac{1}{|S|}\right| \geq \frac{\epsilon |S|}{|S|}\right) \\ &= \Pr\left(\left|\frac{1}{|S|} \sum_i Y_i - \mu\right| \geq \epsilon \mu\right) \end{aligned}$$

~~case~~

$$= \Pr\left(\left|\sum_i Y_i - t\mu\right| \geq \epsilon t\mu\right)$$

$$= \Pr\left(\left|\sum_i Y_i - \hat{\mu}\right| \geq \epsilon \hat{\mu}\right)$$

$$\begin{aligned} \text{where } \hat{\mu} &= t\mu \\ &= E[Y] \\ &= E[\sum_i Y_i] \end{aligned}$$

$$\leq 2e^{-\hat{\mu}\epsilon^2/3} \quad (\text{by Chernoff bounds})$$

$$= 2e^{-\frac{|S|}{|S|}\epsilon^2/3} = 2e^{-t\mu\epsilon^2/3}$$

$$\leq \delta$$

$$\text{for } t \geq \frac{43 \ln(2/\delta)}{\mu \epsilon^2}$$

So only efficient if $\mu = \Omega(\frac{1}{\text{Poly}(n)})$

(5)

Back to #DNF.

Naive approach:

Generate α random assignments X_1, \dots, X_α .

Let $Y_i = \begin{cases} 1 & \text{if assignment } X_i \text{ satisfies } f \\ 0 & \text{otherwise.} \end{cases}$

$$E[Y_i] = \Pr(Y_i = 1) = \frac{N(f)}{2^n}$$

$$\text{Let } Z = \sum_{i=1}^{\alpha} Y_i$$

$$\text{& thus } E[Z] = \sum_{i=1}^{\alpha} \frac{N(f)}{2^n} = \alpha \frac{N(f)}{2^n}$$

$$\text{Note, } \mu = \frac{N(f)}{2^n} \text{ & thus for } t \geq \frac{3 \ln(\frac{2}{\delta})}{\mu e^2} \\ = \frac{\alpha \cdot 3 \ln(\frac{2}{\delta})}{N(f) e^2}$$

we have an (ϵ, δ) -approximation scheme.

But it may be that $\frac{\alpha}{N(f)}$ is huge.

What to do when:

$$N(f) \ll 2^n ?$$

For clauses C_1, \dots, C_m ,

let $S_i = \text{assignments which satisfy } C_i$

& let ~~Σ~~ = $\bigcup_i S_i$

Note, $N(f) = |\Sigma|$.

We can easily sample uniformly at random from S_i

by satisfying the literals in C_i &

then choosing a random assignment for the remaining variables.

Moreover, if k variables appear in C_i (i.e., $|C_i| = k$)

then $|S_i| = 2^{n-k}$.

Consider: $A = \{(i, \sigma) : \sigma \in S_i\}$

which is the multiset union of the S_i 's.

Note, $|A| \leq m \times |\Sigma|$.

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We can sample uniformly at random from 1 & then use that to estimate $|S| = N(f)$.

For $j=1 \rightarrow t$:

1. Choose $i \in \{1, \dots, m\}$ with prob. $\frac{|S_i|}{Z}$

$$\text{where } Z = \sum_i |S_i|.$$

2. Choose σ u.a.r. from S_i

3. Let $Y_j = \frac{Z}{N(\sigma)}$ where $N(\sigma) = \# \text{ of clauses satisfied by } \sigma$.

Output $Y = \frac{1}{t} \sum_{j=1}^t Y_j$

$$\mathbb{E}[Y_j] = \sum_i \frac{|S_i|}{Z} \sum_{\sigma \in S_i} \frac{1}{|S_i|} \frac{z}{N(\sigma)}$$

$$= \sum_i \sum_{\sigma \in S_i} \frac{1}{N(\sigma)} = |S_i| = N(f).$$

$$= \sum_{\sigma \in S} \sum_{i: \sigma \in S_i} \frac{1}{N(\sigma)} = \sum_{\sigma \in S} 1 = |S| = N(f).$$

~~Now by Chernoff bounds,~~

$$\Pr(|N - \mathbb{E}[Y]| \geq \epsilon \mathbb{E}[Y]) \leq 2e^{-\mathbb{E}[Y]\epsilon^2/3}$$

~~Note, Y_j is possibly huge so not clear how to apply Chernoff bounds.~~

⑨

$$E[Y_i^2] = \sum_i \frac{|S_i|}{Z} \sum_{\sigma \in S_i} \frac{1}{|S_i|} \frac{Z^2}{N(\sigma)^2}$$

$$= \sum_i \sum_{\sigma \in S_i} \frac{Z}{N(\sigma)^2}$$

$$= \sum_{\sigma \in S_Z} \sum_{i: \sigma \in S_i} \frac{Z}{N(\sigma)^2}$$

$$= \sum_{\sigma \in S_Z} \frac{Z}{N(\sigma)}$$

$$\leq \sum_{\sigma \in S_Z} Z = Z \sum_{\sigma \in S_Z} 1 = Z \times |S_Z|$$

$$= Z \times N(f).$$

$$\text{Var}(Y_j) = E[Y_j^2] - E[Y_j]^2$$

$$\leq Z \times N(f) - N(f)^2$$

$$= N(f)^2 \left(\frac{Z}{N(f)} - 1 \right) \leq N(f)^2 (m-1)$$

~~because $Z \leq mN(f)$.~~

$$Z = \sum_i \sum_{\sigma \in S_i} 1 = \sum_{\sigma \in S_Z} \sum_{i: \sigma \in S_i} 1$$

$$\leq m \sum_{\sigma \in S_Z} 1$$

$$= m N(f).$$

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Note, $Y = \frac{1}{m} \sum_{j=1}^m Y_j$

Clearly, $E[Y] = E[Y_j] = N(f)$.

$$\& \text{Var}(Y) = \underbrace{\text{Var}(Y_j)}_{+} \leq \frac{(m-1)}{m} N(f)^2$$

~~By Chebyshev's ineq.~~ hence $\sqrt{\text{Var}(Y)} \leq \frac{\sqrt{m-1}}{\sqrt{m}} N(f)$

$$\Pr(|Y - N(f)| \geq \epsilon N(f)) \quad N(f) \geq \sqrt{\frac{\text{Var}(Y) + \epsilon^2 \text{Var}(Y)}{m-1}}$$

$$\leq \Pr\left(|Y - N(f)| \geq \epsilon \sqrt{\frac{\text{Var}(Y) + \epsilon^2 \text{Var}(Y)}{m-1}}\right)$$

$$\leq \Pr\left(|Y - E(Y)| \geq \frac{\epsilon \sqrt{F}}{\sqrt{m-1}} \times \sigma\right) \quad \sigma = \sqrt{\text{Var}(Y)}$$

$$\leq \frac{m-1}{\epsilon^2 F} \leq \frac{1}{4}$$

for ~~ϵ~~ $\epsilon + \frac{1}{\epsilon^2} \geq \frac{4m}{F}$ & then take the median of $O(\log(\delta))$ trials.

(a)

Network unreliability problem:

input: undirected $G = (V, E)$, & parameter $0 \leq p \leq 1$

For each edge $e \in E$, independently delete
with prob. p . Let H be the resulting subgraph.

Let $\text{FAIL}_G(p) = \Pr(\text{resulting } \cancel{\text{subgraph}}^H \text{ is disconnected})$

output: FPRAS for $\text{FAIL}_G(p)$.

For graph H , let $c(H) = \text{min-cut size in } H$ (# of edges)

Note, $\text{FAIL}_G(p) = \Pr(E(G) \setminus E(H) \text{ contains a cut of } G)$
↑
removed a cut so H is disconnected.

Computing $\text{FAIL}_G(p)$ exactly is #P-complete.

Note, $\text{FAIL} = \Pr(E(G) \setminus E(H) \text{ contains a cut})$
 $\geq \Pr(E(G) \setminus E(H) \text{ contains a specific cut})$
 $\geq p^c$

FPRAs due to [Karger '99]:

Naive scheme:

Run the following experiment: l times:

for each edge, delete w.p. p

let $X_i = \begin{cases} 1 & \text{if resulting graph is disconnected} \\ 0 & \text{if connected} \end{cases}$

$$E[X_i] = \text{FAIL}_G(p)$$

$$E[X] = \sum_{i=1}^l X_i \quad \text{for } X = \sum_{i=1}^l X_i$$

assume $p^c \geq n^{-4}$ & thus $\text{FAIL} \geq n^{-4}$.

Then for $l = \frac{3n^4}{\epsilon^2} \log(2/\delta)$, by Chernoff bounds

X is an (ϵ, δ) -approximation of $\text{FAIL}_G(p)$.

What if P^C is tiny?

(C)

Karger's min-cut alg. implies $O(n^2)$ cuts of min-size C .

Moreover, for $\alpha \geq 1$, $O(n^{2\alpha})$ cuts of

Size $\leq \alpha C$.

(Run Karger's alg. down to 2α , instead of 2 vertices.)

And can enumerate all such small cuts in
time $O(n^{2+2\alpha} \log n)$.

Choosing $\alpha = 2 + \ln(\frac{2}{\epsilon})$ & then

cuts larger than αC don't matter

(i.e., have total prob. of failing $\leq \frac{\epsilon}{2}$)

Q
Write a DNF with a variable x_e for each edge $e \in E$.

Make a clause for every cut of size $\leq \Delta C$.

~~Count sat~~

Set each variable to true with prob. p

& false with $1-p$.

What's the probability the resulting formula
is satisfied?

Corresponds to counting $\sum_{\text{Sat. assig. } \sigma} p^{\text{positive}(\sigma)} (1-p)^{\text{negative}(\sigma)}$

where $\text{positive}(\sigma) = \# \text{ of } \cancel{\text{true}} \text{ variables set to}$
true in σ

& $\text{negative}(\sigma) = \# \text{ of } \cancel{\text{variables}} \text{ set to}$
false in σ .

Our alg. for #DNF corresponded to $p = \frac{1}{2}$,
& can easily be generalized to arbitrary p .