

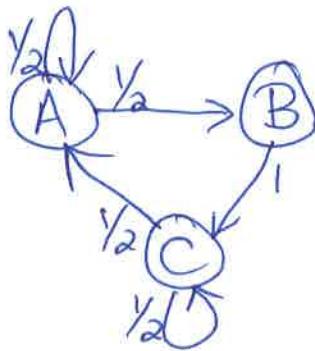
①
For an ergodic MC defined by P on \mathcal{S} with stationary distribution π .

What if we don't know the mixing time?

Can we detect when we've reached π ?

Coupling from the past by [Propp-Wilson '96].

Consider this example with 3 states $\mathcal{S} = \{A, B, C\}$



Idea: Make 3 chains X_+, Y_+, Z_+ where:

$$X_0 = A, Y_0 = B, Z_0 = C.$$

Define a "global" coupling, i.e., a joint evolution of all $|\mathcal{S}|$ chains.

Let T be the 1st time that $X_T = Y_T = Z_T$.

Is this state from distribution π ?
i.e., is $X_T \sim \pi$?

In this example: No it's not from π . ②

Why?

Note $X_T \neq B$.

because we can't ~~1st~~ couple at state B.

If $X_T = Y_T = Z_T = B$ then $X_{T-1} = Y_{T-1} = Z_{T-1} = A$.

But $\pi(B) > 0$ so thus X_T is not from π .

It turns out that if we do this experiment "backwards" it works.

Let's formalize:

Global coupling: for all states $i \in \mathbb{Z}$,

given the "random choice"
it defines the move $i \rightarrow j$.

Think of a transition as choosing $r \in [0, 1]$

uniformly at random,

then the move $i \rightarrow j$ is determined by r.

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Example: Ising model: $\Sigma = \{+1, -1\}^V$, $V = \{0, 1, \dots, n-1\}$

From $X_i \in V$,

1. Choose $r \in [0, 1]$ at random.

2. If $\frac{i}{n} < r \leq \frac{(i+1)}{n}$ then:

a. for all $j \neq i$, set $X_{i+1}(j) = X_i(j)$.

b. Set $X_{i+1}(i) = \begin{cases} +1 & \text{if } n(r - \frac{i}{n}) < \frac{e^{-\beta p}}{e^{-\beta p} + e^{\beta q}} \\ -1 & \text{if } n(r - \frac{i}{n}) \geq \frac{1}{2} \end{cases}$

where $p = \# \text{ of + neighbors of } i \text{ in } X_i$
 $\& q = \# \text{ of - } "$

This is an equivalent form of the Glauber dynamics/
 Gibbs sampler.

Transitions are defined by a function:

$$f: \Sigma \times [0, 1] \rightarrow \Sigma$$

as long as this respects the transition matrix i.e.,

$$\Pr(f(i, r) = j) = P(i, j)$$

Probability is over $r \in [0, 1]$.

④

This also defines a global coupling where all chains use the same random seed $r \in [0, 1]$.

1. Choose $r \in [0, 1]$ at random.

2. If $X_t = i$, set $X_{t+1} = f(i, r)$.

For each time t , choose $r_t \in [0, 1]$ at random.

The transitions for time $t \rightarrow t+1$ are defined by r_t .

For convenience, let $f_t : \mathbb{Z} \rightarrow \mathbb{Z}$ be:

$$f_t = f(\cdot, r_t)$$

$$\text{i.e., } f_t(i) = f(i, r_t).$$

For a chain (X_t) we have that: $X_{t+1} = f_t(X_t)$.

Let $X_0 = i$. What's X_t ?

$$\begin{aligned} \text{Let } F_0^+(i) &= f_{t-1}(f_{t-2}(\dots f_0(i)) \dots) \\ &= (f_{t-1} \circ f_{t-2} \circ \dots \circ f_0)(i) \end{aligned}$$

Then, $X_t = F_0^+(i)$

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More generally, for $0 \leq t_1 < t_2$, let

$$F_{t_1}^{t_2}(i) = (f_{t_2-1} \circ f_{t_2-2} \circ \dots \circ f_{t_1})(i)$$

Hence, for $i, j \in \Sigma$,

$$\Pr(F_{t_1}^{t_2}(i) = j) = P^{t_2-t_1}(i, j).$$

Take the wrong forward algorithm
which finds the \mathbb{P}^T where all chains couple.

We have N chains where $N = |\Sigma|$.

Denote as $(X_1^1), (X_1^2), \dots, (X_1^N)$

where $X_0^i = i$ so i^{th} chain starts @ $i \in \Sigma$.

Run all N chains using f .

Stop when reach time T where $|F_0^T(\Sigma)| = 1$

i.e., for all $i \in \Sigma$, $F_0^T(i) = j$

there exists $j \in \Sigma$, (all chains are in state j at time T)

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Then we output j & hope this is from π ?

The earlier example shows that this is false.

What if we go back in time?

Let M be the first time where

$$|F_m^o(\Sigma)| = 1.$$

Then output $F_m^o(\Sigma)$.

Theorem: $F_m^o(\Sigma)$ has the same distribution as π .

Proof:

For fixed $t > 0$ & $i, j \in \Sigma$, note:

$$\Pr(F_o^+(i) = j) = \Pr(F_{-+}^o(i) = j)$$

↑
π

Why?

This is over r_0, \dots, r_{t-1}

this is over r_{-t}, \dots, r_{-1}

but same distributions, just diff f. names for these random seeds.

Thus, for all $i, j \in \Sigma$,

$$\lim_{t \rightarrow \infty} \Pr(F_{-t}^o(i) = j) = \lim_{t \rightarrow \infty} \Pr(F_t^o(i) = j)$$
$$= \pi(j) \quad \text{since it's ergodic with stationary dist. } \pi$$

For $t = t_1 + t_2$, note:

$$F_{-t}^o = F_{-t_2}^{-1} \circ F_{-t_1}^o$$

Thus if F_{-m}^o is a constant function ($|F_{-m}^o(\Sigma)| = 1$)
then for all $t > M$, all $i \in \Sigma$,

$$F_{-t}^o(i) = (F_{-t}^{-M-1} \circ F_{-m}^o)(i) = F_{-m}^o(i).$$

Hence, $F_{-m}^o(i) \sim \lim_{t \rightarrow \infty} F_{-t}^o(i) \sim \pi$



Intuition?

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Going forward: Let $T = 1^{\text{st}}$ time where F_0^+ is a constant function
for $t > T$ we know F_0^+ is also a constant function
but we don't know that: $F_0^+(i) \stackrel{?}{=} F_T^+(i)$
So we don't know the distribution of $F_0^+(i)$.

Going backwards: for $t > M$:

$$F_{-+}^0(i) = F_{-M}^0(i)$$

So it converges to π .

because $\lim_{t \rightarrow \infty} F_{-+}^0(i) \sim \pi$.

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Can we implement it efficiently?

$|S_2|$ is HUGE so can't run $|S_2|$ chains.

If MC is monotone then we can do it.

Ising model:

For $X_+, Y_+ \in S_2 = \{+1, -1\}^V$

Say $X_+ \geq Y_+$ if for all $v \in V$,

$$X_+(v) \geq Y_+(v),$$

(i.e., if $Y_+(v) = +1$ then $X_+(v) = +1$)

If $X_+ \rightarrow X_{++}$ & $Y_+ \rightarrow Y_{++}$ use same $r \in \{0, 1\}$

then if $X_+ \geq Y_+$ then $X_{++} \geq Y_{++}$.

Why? Both chains update the same vertex.

If update v then

$$\frac{\#\text{ of + neighbors}}{\text{of } v \text{ in } X_+} \geq \frac{\#\text{ of + neighbors}}{\text{of } v \text{ in } Y_+}$$

hence

if $Y_{++}(v) = +1$ then $X_{++}(v) = +1$.

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Consider $W_0 = \text{all} +$ & $Y_0 = \text{all} -$

& arbitrary $X_0 \in \Sigma$.

Then, $W_0 \geq X_0 \geq Y_0$

& for all $t \geq 0$,

$$W_t \geq X_t \geq Y_t$$

Take t where $W_t = Y_t$ then

for all X_0 , $W_t = X_t = Y_t$

$$\text{So } |F_0^+(S)| = 1.$$

Hence, run W_t & Y_t and

look for min M so that $W_M = Z_M$

& then we have that $|F_M^+(S)| = 1$

So we're done.

So for monotone system just need to consider 2 chains, instead of $|\Sigma|$ chains.

What is M ?

Is it much larger than the mixing time?

For Monotone MCS,

$$E[M] \leq 2T_{\text{mix}} \ln(4n).$$

So it's not much more than the mixing time.